# Studying Star Formation through Hierarchical Bayesian Modeling of Emission from Astronomical Dust

Brandon C. Kelly (CfA), Rahul Shetty (ITA, Heidelberg), Amelia M. Stutz (MPIA), Jens Kauffmann (JPL), Alyssa Goodman (CfA), Ralf Launhardt (MPIA)

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#### Using Dust to Study Star Formation







Image Credit: NASA/JPL

Astrostatistics Class 10/6/09

# Thermal Emission

- Objects with a temperature T emit thermal emission
  - Idealized spectrum of thermal emission is the Planck function
  - Often called `blackbody emission'
  - Cooler object emit more energy at longer wavelengths
- Dust is an efficient radiator of thermal energy
- For dust between stars, emission is dominated by Far-IR to sub-mm wavelengths.



Mode and Normalization of BB increase as temperature increases



# **Modeling Dust Emission**

Model dust brightness as a modified black-body:

$$\left| f(\mathbf{v}) \propto N \mathbf{v}^{\beta} B_{\mathbf{v}}(T), \ B_{\mathbf{v}}(T) \propto \mathbf{v}^{3} \left[ \exp \left( \frac{h \mathbf{v}}{kT} \right) - 1 \right]^{-1} \right|$$

- Parameters are the dust temperature, T, the power-law modification index, β, and the column density, N
- β -> 0 as dust grains become larger





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# Astrophysical Expectations

- Dust coagulation: Higher dust agglomeration in dense regions should create larger grains
  - β should decrease toward denser regions

- For dust between stars with silicate and graphite composition,  $\beta\approx 2$
- For disks of dust around new stars, observations find  $\beta \le 1$ .
- Temperature should decrease in higher density regions
  - Higher density regions more effectively shielded from ambient radiation field

## Previous Observational Results



Observations find β and temperature anti-correlated (e.g., Dupac+2003, Desert+2008, Anderson+2010, Paradis+2010, Planck Collaboration 2011)





## Unexpected: Why a $\beta$ -T anticorrelation?

- β and T estimated by minimizing weighted squared error (i.e., χ<sup>2</sup>)
- Typically only 5-10 data points for estimating 3 parameters
- Errors on estimated β and T large, highly anti-correlated
- Errors bias inferred correlation, may lead to spurious anticorrelation (Shetty+2009)



## The Measurement Model



 $\begin{array}{l} y_{ij}: \text{Measured flux value at } j^{th} \text{ wavelength for } i^{th} \text{ data point} \\ i = 1, \ldots, n \quad j = 1, \ldots, p \\ N_i, T_i, \beta_i: \text{model parameters for } i^{th} \text{ data point} \\ v_j: \text{Frequency corresponding to } j^{th} \text{ observational wavelength} \\ \delta_j = \text{Calibration error for } j^{th} \text{ observational wavelength} \\ \varepsilon_{ij} = \text{Measurement noise for } y_{ij} \end{array}$ 

## Our Hierarchical Model

Joint distribution of log N, log T, and  $\beta$  modeled as a multivariate Normal distribution

$$\mu, \Sigma \sim Uniform$$

$$\log N_i, \log T_i, \beta_i \mid \mu, \Sigma \sim N(\mu, \Sigma)$$

$$\log \delta_j \sim N(0, \tau_j^2)$$

$$y_{ij} \mid N_i, T_i, \beta_i, \delta_j \sim N(\delta_j S_j(N_i, T_i, \beta_i), \sigma_{ij}^2)$$

 $\tau_i$  and  $\sigma_{ii}$  are considered known and fixed

Calibration Errors are nuisance parameters

## Naïve MCMC Sampler

 Update calibration error, do Metropolis-Hastings (M-H) update with proposal

$$\begin{split} & \left[ \log \delta_{j}^{\text{prop}} \sim N(\log \overline{\delta}_{j}, Var(\overline{\delta}_{j}) / \overline{\delta}_{j}^{2}) \\ & \overline{\delta}_{j} = Var(\overline{\delta}_{j}) \left( \frac{\overline{\delta}_{j}^{\text{data}}}{v_{j}^{\text{data}}} + \frac{1}{\tau_{j}^{2}} \right), \quad Var(\overline{\delta}_{j}) = \left( \frac{1}{v_{j}^{\text{data}}} + \frac{1}{\tau_{j}^{2}} \right)^{-1} \\ & \overline{\delta}_{j}^{\text{data}} = v_{j}^{\text{data}} \sum_{i=1}^{n} \frac{y_{ij}}{S_{j}(N_{i}, T_{i}, \beta_{i})} \left( \frac{\sigma_{ij}}{S_{j}(N_{i}, T_{i}, \beta_{i})} \right)^{-2}, \quad v_{j}^{\text{data}} = \left( \sum_{i=1}^{n} \left( \frac{\sigma_{ij}}{S_{j}(N_{i}, T_{i}, \beta_{i})} \right)^{-2} \right)^{-1} \end{split}$$

- 2. Update values of  $N_i$ ,  $T_i$ , and  $\beta_i$  using M–H update with multivariate normal proposal density
- 3. Do Gibbs update of  $\mu$  and  $\Sigma$  using standard updates for mean and covariance matrix of Normal distribution



## Ancillary-Sufficiency Interweaving Strategy

 Naïve MCMC sampler is very slow due to strong dependence between δ, log N, β, and μ:

$$E(y_{ij} | N_i, T_i, \beta_i, \delta_j) \propto \delta_j N_i \left(\frac{\nu_j}{\nu_0}\right)^{\beta_i} B_{\nu_j}(T_i)$$

 Use Ancillary-Sufficiency Interweaving Strategy (ASIS, Yu & Meng 2011) to break dependence

References:

Yu, Y, & Meng, X–L, 2011, JCGS (with discussion), 20, 531

Kelly, B.C., 2011, JCGS, 20, 584

#### Introduce Sufficient Augmentation

- Original data augmentation is an *ancillary* augmentation for calibration error: δ and (β,N,T) independent in their prior
- Introduce a *sufficient* augmentation for calibration error:

$$\boxed{ \begin{split} &\log \tilde{\delta}_i = \log \delta + \mathbf{X} \theta_i \\ & \left( \begin{array}{cc} 1 & \log \frac{\nu_1}{\nu_0} \\ \vdots & \vdots \\ 1 & \log \frac{\nu_p}{\nu_0} \end{array} \right), \quad \theta_i = (\log N_i, \beta_i)^T \end{split} }$$

## Sufficient Augmentation Model

New Model is

$$\begin{aligned} y_{ij} \mid T_i, \tilde{\delta}_{ij} \sim N(\tilde{\delta}_{ij}B_j(T_i), \sigma_{ij}^2) \\ \log \tilde{\delta}_i \mid \log \tilde{\delta}_k, \theta_i, \theta_k &= \log \tilde{\delta}_k - \mathbf{X}(\theta_k - \theta_i), i \neq k \\ \log \tilde{\delta}_k \mid \theta_k \sim N(\mathbf{X}\theta_k, V_\delta), \quad V_\delta &= \operatorname{diag}(\tau_j^2) \\ (\theta_i, \log T_i) \mid \mu, \Sigma \sim N(\mu, \Sigma) \end{aligned}$$



# ASIS for this problem

- 1. Update calibration error as before
- 2. Update  $\delta$ , log N, and  $\beta$  under SA:

$$\begin{aligned} & \theta_{k}^{new} \mid \tilde{\delta}_{k} \sim N(\hat{\theta}_{k}, V_{k}) \\ & \hat{\theta}_{k} = V_{k} \mathbf{X}^{T} V_{\delta}^{-1} \tilde{\delta}_{k} \\ & V_{k} = (\mathbf{X}^{T} V_{\delta}^{-1} \mathbf{X})^{-1} \\ & \theta_{i}^{new} \mid \theta_{k}^{new}, \tilde{\delta}_{k}, \tilde{\delta}_{i} = \theta_{k}^{new} + (\mathbf{X}^{T} \mathbf{X})^{-1} \mathbf{X}^{T} (\tilde{\delta}_{i} - \tilde{\delta}_{k}) \\ & \delta = \tilde{\delta}_{k} - \mathbf{X} \theta_{k}^{new} \end{aligned}$$

(also need to update µ appropriately)

- 3. Update log N, log T, and  $\beta$  as before under AA
- 4. Update  $\mu$  and  $\Sigma$

## ASIS Significantly Improves MCMC Efficiency for this Problem



MCMC results for data point with highest S/N, 200,000 iterations

Kelly 2011

# Test: Comparison of Bayesian and $\chi^2$ methods for Intrinsic Correlation



Rank correlation coefficients: True = 0.33

 $\begin{array}{l} \chi^2 = -0.45 {\pm} 0.03 \\ Bayes = 0.23 {\pm} 0.08 \end{array}$ 

## Application: Star-Forming Bok Globule CB244

CB244: Small, nearby molecular cloud containing a low-mass protostar and starless core



## CB244 Results:

- β and T anti-correlated for χ<sup>2</sup>-based estimates, caused by noise
- β and T *correlated* for Bayesian estimates, opposite what has been seen in previous work

Random Draw from Posterior Probability Distribution



#### CB244: Temperature and Column Density Maps





Temperature tends to decrease toward central, denser regions

### CB244: Column Density and β Maps





β decreases toward central denser regions

# Evidence for Grain Growth?

- β is expected to be smaller for larger grains
- β < 1 for protoplanetary disks, denser than starless cores
- For CB244 we find β is smaller in denser regions
- Suggests dust grains begin growing on large scales before star forms
- Qualitatively consistent with spectroscopic features seen in mid-IR observations of starless cores (Stutz+2009)



# Summary

- Introduced a Bayesian method that correctly recovers intrinsic correlations involving N, $\beta$ , and T, in contrast to traditional  $\chi^2$
- Developed an ancillarity-sufficiency interweaving strategy to make problem computationally tractable
- > For CB244, our Bayesian method finds that  $\beta$  and T are correlated, opposite that of  $\chi^2$
- For CB244, we also find that  $\beta$  increases toward the central dense regions, while temperature decreases
- > Increase in  $\beta$  with column density may be due to growth of dust grains through coagulation
- Bayesian method led to very different scientific conclusions compared to naïve χ<sup>2</sup> method, illustrating importance of proper statistical modeling for complex astronomical data sets

